ECON 3740: INTRODUCTION TO ECONOMETRICS

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Lecture 4

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Lecture outline

Last lecture, we learned the types of data, some terminologies for the simple linear regression model, and part of the derivation of OLS estimators. Today, we will

- complete the derivation for the OLS estimators
 - solve for the $\hat{\beta}_0$ & $\hat{\beta}_1$ based on the two FOCs we derived in last lecture
- introduce some measures of fit
 - SST
 - SSE
 - SSR
 - R²
- change the units measurement
- handle the technique to characterize non-linearity
 - take logs of the dependent variable only
 - take logs of the independent variable only
 - take logs of both

The ordinary least squares estimator (OLS): derivation

• Now, we have two FOCs.

$$1/n\sum_{i=1}^{n} \left(y_i - \hat{\beta}_0 - \hat{\beta}_1 x_i \right) = 0$$

$$1/n \sum_{i=1}^{n} x_i \left(y_i - \hat{\beta}_0 - \hat{\beta}_1 x_i \right) = 0$$

- We will use these two FOCs to derive β_0 and β_1 .
- Step 1: Derive OLS estimator of β_0 as a function of OLS estimator of β_1

$$\frac{1/n\sum_{i=1}^{n} \left(y_{i} - \hat{\beta}_{0} - \hat{\beta}_{1} x_{i} \right) = 0}{\Longrightarrow 1/n\sum_{i=1}^{n} y_{i} - 1/n\sum_{i=1}^{n} \hat{\beta}_{0} - 1/n\sum_{i=1}^{n} \hat{\beta}_{1} x_{i} = 0}{\Longrightarrow 1/n\sum_{i=1}^{n} y_{i} - \frac{1}{n}n\hat{\beta}_{0} - \hat{\beta}_{1}\frac{1}{n}\sum_{i=1}^{n} x_{i} = 0}{\Longrightarrow \bar{y}_{i} - \hat{\beta}_{0} - \hat{\beta}_{1}\bar{x}_{i} = 0}$$

The last equation gives

$$\hat{\beta}_0 = \bar{y}_i - \hat{\beta}_1 \bar{x}_i$$

The ordinary least squares estimator (OLS): derivation

• Step 2: Solve for the OLS estimator of β_1 $(\hat{\beta}_1)$

$$\sum_{i=1}^{n} x_i \left(y_i - \hat{\beta}_0 - \hat{\beta}_1 x_i \right) = 0$$

$$\implies \sum_{i=1}^{n} x_i \left(y_i - \left(\bar{y}_i - \hat{\beta}_1 \bar{x}_i \right) - \hat{\beta}_1 x_i \right) = 0$$

$$\implies \sum_{i=1}^{n} x_i \left(\left(y_i - \bar{y}_i \right) - \left(\hat{\beta}_1 x_i - \hat{\beta}_1 \bar{x}_i \right) \right) = 0$$

$$\implies \sum_{i=1}^{n} x_i (y_i - \bar{y}_i) - \hat{\beta}_1 \sum_{i=1}^{n} x_i (x_i - \bar{x}_i) = 0$$

Simple Algebra trick can show

$$\implies \sum_{i=1}^{n} \left(x_i - \bar{x}_i \right) \left(y_i - \bar{y}_i \right) - \hat{\beta}_1 \sum_{i=1}^{n} \left(x_i - \bar{x}_i \right) \left(x_i - \bar{x}_i \right) = 0$$

The ordinary least squares estimator (OLS): derivation

$$\implies \hat{\beta}_1 = \frac{\sum_{i=1}^n \left(x_i - \bar{x}_i\right) \left(y_i - \bar{y}_i\right)}{\sum_{i=1}^n \left(x_i - \bar{x}_i\right) \left(x_i - \bar{x}_i\right)} = \frac{S_{xy}}{S_x^2}$$

- Then, $\hat{\beta}_0$ can be calculated by substituting above equation into $\hat{\beta}_0 = \bar{y}_i \hat{\beta}_1 \bar{x}_i$.
- The OLS predicted values are

$$\hat{y}_i = \hat{\beta}_0 + \hat{\beta}_1 x_i$$

The OLS residuals are

$$\hat{\mu}_i = y_i - \hat{y}_i$$

The ordinary least squares estimator (OLS): measures of fit

- How well does the estimated regression line describe the data?
 - Does the regressor X account for much or for little variation in Y ?
 - Are the observations in the scatter plot clustered closely around the regression line?
- Some measures can help us answer above questions!
- **SST**: Total sum of squares measures the total amount of variability in the dependent variable. That is, it measures how spread out the *y_i* are in the sample.

$$SST = \sum_{i=1}^{n} (y_i - \bar{y})^2$$

• **SSE**: Explained sum of squares measures the sample variation in the \hat{y}_i (where we use the fact that the mean of $\hat{y}_i = \bar{y}$)

$$SSE = \sum_{i=1}^{n} (\hat{y}_i - \bar{y})^2$$

The ordinary least squares estimator (OLS): measures of fit

• **SSR**: Sum of squared residuals measures the total amount of variability that the model does *not* explain

$$SSR = \sum_{i=1}^{n} (\hat{\mu}_i)^2$$

• R^2 : measures the variation "explained" by the model

$$R^2 = \frac{ESS}{TSS}$$

• The R^2 ranges from 0 to 1

- If $R^2 = 0$, x_i explains zero variation in y_i
- If $R^2 = 1$, x_i explains all of the variation in y_i
- in practice $R^2 \in (0, 1)$

The ordinary least squares estimator (OLS): The R^2

- Relationship among *SST*, *SSR*, *SSE*?
- The total variation in y can always be expressed as the sum of the explained variation and the unexplained variation SSR. Thus

$$SST = SSE + SSR$$
$$\sum_{i=1}^{n} (y_i - \bar{y})^2 = \sum_{i=1}^{n} (\hat{y}_i - \bar{y})^2 + \sum_{i=1}^{n} (\hat{\mu}_i)^2$$

This implies

$$R^{2} = \frac{SSE}{SSR} = \frac{SST - SSR}{SST} = 1 - \frac{SSR}{SST}$$

• At home: prove SST = SSE + SSR mathematically.

The ordinary least squares estimator (OLS): An example for R^2

• Recall the wage problem, our model is

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wage = \beta_0 + \beta_1 educ + \mu
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• Once we estimated the model using R, we have following results

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Call.
lm(formula = wage \sim educ)
Residuals:
    Min
             10 Median
                             30
                                    Max
-5 3396 -2 1501 -0 9674 1 1921 16 6085
Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept) -0.90485
                        0.68497
                                 -1.321
                                           0.187
                     0 05325 10 167
                                        <2e-16 ***
educ
             0 54136
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
Residual standard error: 3.378 on 524 degrees of freedom
Multiple R-squared: 0.1648, Adjusted R-squared: 0.1632
F-statistic: 103.4 on 1 and 524 DF, p-value: < 2.2e-16
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The ordinary least squares estimator (OLS): an example for R^2

• Therefore,

 $w\hat{a}ge = -0.90485 + 0.54136 \ educ \ \& R^2 = 0.1648$

- The education level explains 16.5% of the variation in wage based on our sample.
- Even the R^2 in above example is high, generally speaking, in the social sciences, we have to notice that low R^2 in regression equations are not uncommon, especially for cross-sectional analysis. We will discuss this issue more generally under multiple regression analysis. However, it is worth emphasizing now that a seemingly low R^2 does **not** necessarily mean that an OLS regression equation is useless.

The ordinary least squares estimator (OLS): changing units of measurement

- Data Scaling
 - Predictions in different units
 - Different interpretations
- Example:

wage= $\beta_0 + \beta_1 educ + \mu$ educ is in years wage is in dollar

Estimates:

wâge
$$= \hat{eta}_0 + \hat{eta}_1$$
educ

• Intuitively, the result implies one more year education increases $\hat{\beta}_1$ dollars in wage.

The ordinary least squares estimator (OLS): changing units of measurement

• Now, suppose the wage unit changes from dollars to cents,

$$\implies w \hat{a} g e_{dollars} = \frac{1}{100} w \hat{a} g e_{cents} \tag{1}$$

- $\bullet~$ Recall, the original estimation gives $w \hat{a} g e = \hat{\beta}_0 + \hat{\beta}_1 e duc$
- next, we substitute equation (1) in to (2) $\frac{1}{100} w \hat{a} g e_{cents} = \hat{\beta}_0 + \hat{\beta}_1 e duc \qquad (3)$ $w \hat{a} g e_{cents} = 100 \hat{\beta}_0 + 100 \hat{\beta}_1 e duc$
- Intuitively, the result implies one more year education increases $100\hat{\beta}_1$ cents in wage.

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The ordinary least squares estimator (OLS): handling non-linearity

- In the real life, you will often encounter regression equations where the dependent variable appears in logarithmic form. Why we need it? Recall the wage-education case, where we regressed dollars of wage on years of education. Because of the linear nature, $\hat{\beta}_1$ dollars is the increase for either the first year of education or the tenth year of education. This can be highly doubted!
- Ideally, a better characterization of how wage changes with education is that each year of education increases wage by a constant *percentage*.
- Three ways to handle non-linearity
 - Take logs of the dependent variable only
 - Take logs of the independent variable only
 - Take logs of both

The ordinary least squares estimator (OLS): handling non-linearity

• Case 1: wage is in logs (take logs of the dependent variable only) $log(w\hat{a}ge) = \hat{\beta}_0 + \hat{\beta}_1 educ$

- How do we interpret $\hat{\beta}_1$?
- Totally differentiate gives

$$rac{1}{w \hat{a} g e} \partial w \hat{a} g e = \hat{eta}_1 \partial e d u c$$

• To simplfy, we have
$$\underbrace{\frac{\partial w \hat{a} g e}{w \hat{a} g e} \times 100}_{\% \ change} = (\hat{\beta}_1 \times 100) \underbrace{\frac{\partial e d u c}{\partial e d u c}}_{unit \ change}$$

• Therefore, we can interpret the story as a one-year increase in education tields a $\hat{\beta}_1$ % increase in wage

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